

Random vectors II.

(Covariance and correlation, convolution)

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Probability and Statistics

BIE-PST, WS 2025/26, Lecture 7



Content

- **Probability theory:**

- ▶ Events, probability, conditional probability, Bayes' Theorem, independence of events.
- ▶ Random variables, distribution function, functions of random variables, characteristics of random variables: expected value, variance, moments, generating function, quantiles, critical values, important discrete and continuous distributions.
- ▶ Random vectors, joint and marginal distributions, independence of random variables, conditional distribution, **functions of random vectors, covariance and correlation**.
- ▶ Markov's and Chebyshev's inequality, weak law of large numbers, strong law of large numbers, Central limit theorem.

- **Mathematical statistics:**

- ▶ Point estimators, sample mean, sample variance, properties of point estimators, Maximum likelihood method.
- ▶ Interval estimators, hypothesis testing, one-sided vs. two-sided alternatives, linear regression, estimators of regression parameters, testing of linear model.

Recap

Joint distribution function of a random vector (X, Y) :

$$F_{X,Y}(x, y) = P(X \leq x \cap Y \leq y).$$

Discrete random variables X and Y

Joint probabilities of values:

$$P(X = x \cap Y = y)$$

Continuous random variables X and Y

Joint density:

$$f_{X,Y}(x, y)$$

Marginal distributions:

$$P(X = x) = \sum_{\text{all } y} P(X = x \cap Y = y)$$

$$P(Y = y) = \sum_{\text{all } x} P(X = x \cap Y = y)$$

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy$$

$$f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dx$$

Independence of X and Y :

$$P(X = x \cap Y = y) = P(X = x) P(Y = y) \quad | \quad f_{X,Y}(x, y) = f_X(x) f_Y(y)$$

Conditional probabilities / density of X given $Y = y$:

$$P(X = x|Y = y) = \frac{P(X = x \cap Y = y)}{P(Y = y)}$$

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x, y)}{f_Y(y)}$$

Conditional expectation of X given $Y = y$:

$$E(X|Y = y) = \sum_x x P(X = x|Y = y) \quad | \quad E(X|Y = y) = \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx$$

Functions of random vectors of variables

Similar formulas as for a function of one random variable also hold for the functions of random vectors. Let

$$Z = h(X_1, \dots, X_n) = h(\mathbf{X}).$$

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- When variables $X_1, \dots, X_n$ have a **joint discrete** distribution with probabilities $P(\mathbf{X} = \mathbf{x})$, the following relation holds for the distribution function of Z :

$$F_Z(z) = P(Z \leq z) = \sum_{\{\mathbf{x} \in \mathbb{R}^n : h(\mathbf{x}) \leq z\}} P(\mathbf{X} = \mathbf{x}).$$

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- When variables $X_1, \dots, X_n$ have a **joint continuous** distribution with density $f_{\mathbf{X}}(\mathbf{x})$, the distribution function of Z is then

$$F_Z(z) = P(Z \leq z) = \int \cdots \int_{\{\mathbf{x} \in \mathbb{R}^n: h(\mathbf{x}) \leq z\}} f_{\mathbf{X}}(\mathbf{x}) \, dx_1 \dots dx_n.$$

Expected value of the function of a random vector

The expected value $E h(X, Y)$ of a real function h of random variables X and Y can be computed without determining the distribution of the variable $h(X, Y)$.

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$$E h(X, Y) = \sum_{i,j} h(x_i, y_j) P(X = x_i \cap Y = y_j),$$

if the sum converges absolutely.

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- For X and Y continuous random variables it holds that

$$E h(X, Y) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} h(x, y) f_{X,Y}(x, y) dx dy,$$

if the integral converges absolutely.

Properties of the expected value

Now we can prove the linearity of the expectation.

Theorem – linearity of expectation

For all $a, b \in \mathbb{R}$ and all random variables X and Y it holds that

$$\mathbb{E}(aX + bY) = a \mathbb{E} X + b \mathbb{E} Y.$$

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Consequence:

- $\mathbb{E}(aX + b) = a \mathbb{E} X + b$. This statement was proven before separately.

Properties of the expected value

Proof

From the theory concerning the marginal distributions of discrete random variables X and Y we have:

$$\begin{aligned} E(aX + bY) &= \sum_{i,j} (ax_i + by_j) P(X = x_i \cap Y = y_j) \\ &= \sum_{i,j} ax_i P(X = x_i \cap Y = y_j) + \sum_{i,j} by_j P(X = x_i \cap Y = y_j) \\ &= a \sum_i x_i \sum_j P(X = x_i \cap Y = y_j) + b \sum_j y_j \sum_i P(X = x_i \cap Y = y_j) \\ &= a \sum_i x_i P(X = x_i) + b \sum_j y_j P(Y = y_j) = a E X + b E Y. \end{aligned}$$

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 \end{aligned}$$

For continuous X and Y the proof is analogous:

$$\begin{aligned}
 E(aX + bY) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (ax + by) f_{X,Y}(x, y) \, dx \, dy = \dots = \\
 &= a \int_{-\infty}^{\infty} x f_X(x) \, dx + b \int_{-\infty}^{\infty} y f_Y(y) \, dy = a E X + b E Y.
 \end{aligned}$$



Covariance and correlation coefficient

Mutual linear dependence of two random variables X and Y can be described in the following way:

Definition

Let X and Y be random variables with finite second moments. Then we define the **covariance** of the random variables X and Y as

$$\text{cov}(X, Y) = E[(X - E X)(Y - E Y)].$$

If X and Y have positive variances then we define the **correlation coefficient** (or **coefficient of correlation**) as

$$\rho(X, Y) = \frac{\text{cov}(X, Y)}{\sqrt{\text{var } X} \sqrt{\text{var } Y}}.$$

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Definition

Two random variables X and Y are called **non-correlated** if $\text{cov}(X, Y) = 0$.

Covariance and the correlation coefficient – properties

Theorem

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i) $\text{cov}(X, Y) = E XY - E X E Y,$

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- iii) $\rho(X, Y) \in [-1, 1]$,
- iv) $\rho(aX + b, cY + d) = \rho(X, Y)$ for all $a, c > 0$ and $b, d \in \mathbb{R}$,

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- v) $\rho(X, Y) = \pm 1$, if $a, b \in \mathbb{R}$, $a > 0$ such that $Y = \pm aX + b$.

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Proof

- i)
$$\begin{aligned} \text{cov}(X, Y) &= E((X - E X)(Y - E Y)) = E(XY - X E Y - Y E X + E X E Y) \\ &= E XY - E(X E Y) - E(Y E X) + E(E X E Y) \\ &= E XY - E X E Y - E Y E X + E X E Y = E XY - E X E Y \end{aligned}$$
- ii) Obvious from above.
- iii) From the Schwarz inequality (see bibliography).
- iv) Follows straightforwardly by inserting into the definition.
- v) Follows from the proof of the Schwarz inequality (see bibliography).



Non-correlated random variables

Let us study the **expectation of the product** XY of two random variables X and Y .

Definition

Alternative definition: Two random variables X and Y are called **non-correlated** if

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Lemma

If X and Y are independent then they are non-correlated.

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Lemma

If X and Y are independent then they are non-correlated.

Proof

Let X, Y be continuous variables. Independence means that $f_{X,Y}(x, y) = f_X(x)f_Y(y)$. Thus we have

$$\begin{aligned} E XY &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xy f_{X,Y}(x, y) \, dx \, dy = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xy f_X(x) f_Y(y) \, dx \, dy \\ &= \left(\int_{-\infty}^{+\infty} x f_X(x) \, dx \right) \left(\int_{-\infty}^{+\infty} y f_Y(y) \, dy \right) = E X E Y. \end{aligned}$$



Properties of the variance

It is now possible to obtain the following properties of the variance of sums of two random variables.

Theorem

i) For X and Y with finite second moments:

$$\text{var}(X \pm Y) = \text{var } X + \text{var } Y \pm 2 \text{cov}(X, Y).$$

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i) For X and Y with finite second moments:

$$\text{var}(X \pm Y) = \text{var } X + \text{var } Y \pm 2 \text{cov}(X, Y).$$

ii) For non-correlated (independent) random variables it holds that

$$\text{var}(X \pm Y) = \text{var } X + \text{var } Y.$$

Properties of variance

Proof

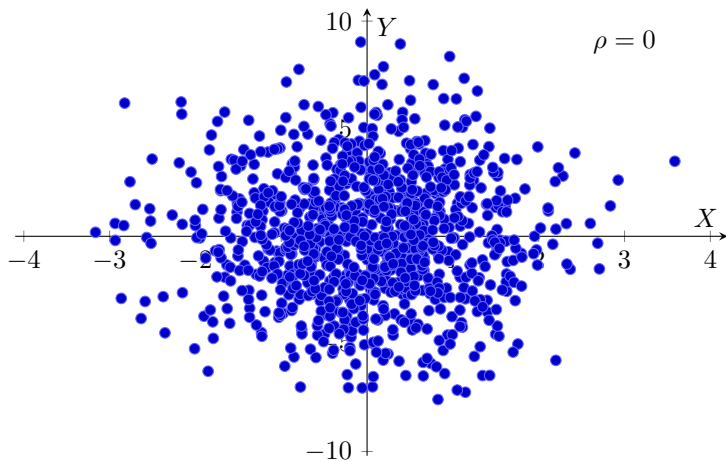
i) Given two random variables X and Y we have:

$$\begin{aligned}\text{var}(X \pm Y) &= E(X \pm Y)^2 - (E(X \pm Y))^2 = E(X^2 \pm 2XY + Y^2) - (E X \pm E Y)^2 \\ &= E X^2 \pm 2 E X Y + E Y^2 - (E X)^2 \mp 2 E X E Y - (E Y)^2 \\ &= \text{var } X + \text{var } Y \pm (2 E X Y - 2 E X E Y) = \text{var } X + \text{var } Y \pm 2 \text{cov}(X, Y).\end{aligned}$$

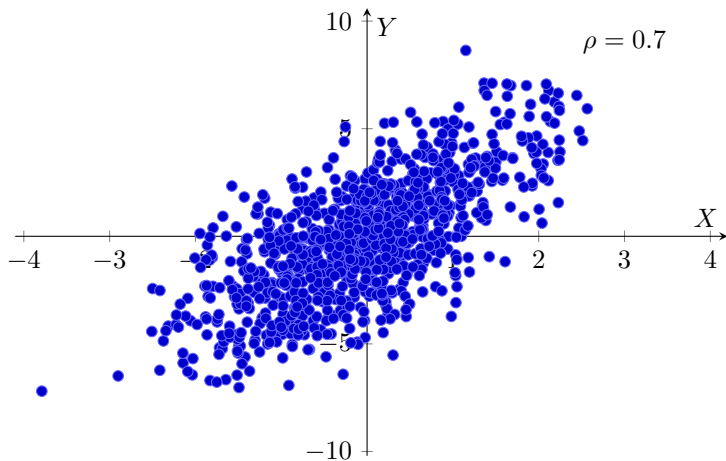
ii) For non-correlated (independent) random variables the covariance is zero.



Correlation – sample of 1000 values



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Sums of random variables

An important case of a function of multiple random variables is their sum

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Consider for simplicity a sum of two random variables:

- If X and Y are **discrete** and **independent**, then for $Z = X + Y$ it holds that

$$P(Z = z) = \sum_x P(X = x) \cdot P(Y = z - x) \quad (\text{discrete convolution}).$$

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- If X and Y are **continuous** and **independent**, then for $Z = X + Y$ it holds that

$$f_Z(z) = \int_{-\infty}^{\infty} f_X(x) f_Y(z - x) \, dx \quad (\text{convolution of } f_X \text{ and } f_Y).$$

Sums of random variables – convolution (discrete case)

The expression for the sum of discrete independent X and Y is obtained easily:

$$\begin{aligned} P(Z = z) &= P(X + Y = z) \\ &= \sum_{\{(x_k, y_j): x_k + y_j = z\}} P(X = x_k \cap Y = y_j) \\ &= \sum_{\text{all } x_k} P(X = x_k) P(Y = z - x_k). \end{aligned}$$

Sums of random variables - convolution (continuous case)

For continuous independent X and Y we have:

$$\begin{aligned}
 F_Z(z) = P(X + Y \leq z) &= \iint_{\{(x,y): x+y \leq z\}} f_{X,Y}(x,y) \, d(x,y) \\
 &= \int_{-\infty}^{\infty} \left(\int_{-\infty}^{z-x} f_{X,Y}(x,y) \, dy \right) dx \\
 &\stackrel{y=u-x}{=} \int_{-\infty}^{\infty} \left(\int_{-\infty}^z f_{X,Y}(x, u-x) \, du \right) dx \\
 &= \int_{-\infty}^z \left(\int_{-\infty}^{\infty} f_{X,Y}(x, u-x) \, dx \right) du \\
 &= \int_{-\infty}^z \left(\int_{-\infty}^{\infty} f_X(x) f_Y(u-x) \, dx \right) du.
 \end{aligned}$$

The density f_Z is any non-negative function, for which $F_Z(z) = \int_{-\infty}^z f_Z(u) \, du$.

The expression under the first integral $f_Z(z) = \int_{-\infty}^{\infty} f_X(x) f_Y(z-x) \, dx$ is thus the density of Z .

Sum of random variables – Poisson distribution

Example

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$$P(X = j) = \frac{\lambda_1^j}{j!} e^{-\lambda_1} \quad P(Y = \ell) = \frac{\lambda_2^\ell}{\ell!} e^{-\lambda_2}, \quad j, \ell = 0, 1, \dots$$

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From what we have seen before we know that for $k = 0, 1, \dots$:

$$P(Z = k) = \sum_{\{(j, \ell) \in \mathbb{N}_0^2 : j + \ell = k\}} P(X = j) P(Y = \ell) = \sum_{i=0}^k P(X = j) P(Y = k - j)$$

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Sum of random variables – Normal distribution

Example – sum of two normal distributions

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The densities of X and Y correspond to the normal distribution with variance $\sigma^2 = 1$:

$$f_X(x) = \frac{1}{\sqrt{2\pi \cdot 1}} e^{-\frac{(x-\mu)^2}{2 \cdot 1}}, \quad f_Y(y) = \frac{1}{\sqrt{2\pi \cdot 1}} e^{-\frac{(y-\mu)^2}{2 \cdot 1}} \quad x, y \in \mathbb{R}.$$

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The density of the sum is obtained using convolution:

$$\begin{aligned} f_Z(z) &= \int_{-\infty}^{\infty} f_X(x) f_Y(z-x) dx = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2}} \frac{1}{\sqrt{2\pi}} e^{-\frac{(z-x-\mu)^2}{2}} dx \\ &= \int_{-\infty}^{\infty} \frac{1}{2\pi} e^{-\frac{1}{2}((x-\mu)^2 + (z-x-\mu)^2)} dx. \end{aligned}$$

Sum of random variables – Normal distribution

Example – sum of two normal distributions, continuation

The expressions in the exponent can be rewritten as:

$$\begin{aligned}(x - \mu)^2 + (z - x - \mu)^2 &= x^2 - 2\mu x + \mu^2 + z^2 + x^2 + \mu^2 - 2zx - 2\mu z + 2\mu x \\ &= 2 \left(x - \frac{z}{2} \right)^2 + \frac{1}{2} (z - 2\mu)^2.\end{aligned}$$

The expression under the integral can then be split into two multiplicative parts, with one of them not depending on x and the other one having an integral of 1:

$$\begin{aligned}f_Z(z) &= \int_{-\infty}^{\infty} \frac{1}{2\pi} e^{-\frac{2(x-z/2)^2}{2}} e^{-\frac{(z-2\mu)^2}{2 \cdot 2}} dx \\ &= \frac{1}{\sqrt{2\pi}2} e^{-\frac{(z-2\mu)^2}{2 \cdot 2}} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}(1/2)} e^{-\frac{(x-z/2)^2}{2 \cdot (1/2)}} dx \\ &= \frac{1}{\sqrt{2\pi}2} e^{-\frac{(z-2\mu)^2}{2 \cdot 2}}.\end{aligned}$$

The sum $Z = X + Y$ has therefore the normal distribution $N(2\mu, 2)$. In general, it can be proven that the sum of n independent normals $N(\mu, \sigma^2)$ has the distribution $N(n\mu, n\sigma^2)$.

Sums of random variables – moment generating function

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$$\begin{aligned} M_Z(s) &= E(e^{sZ}) = E(e^{s(X+Y)}) = E(e^{sX}e^{sY}) \\ &= E(e^{sX})E(e^{sY}) = M_X(s)M_Y(s). \end{aligned}$$

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Generally for a vector of independent random variables $X_1, \dots, X_n$ it holds that:

$$Z = X_1 + \dots + X_n \implies M_Z(s) = M_{X_1}(s) \cdots M_{X_n}(s).$$

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Generally for a vector of independent random variables $X_1, \dots, X_n$ it holds that:

$$Z = X_1 + \dots + X_n \implies M_Z(s) = M_{X_1}(s) \cdots M_{X_n}(s).$$

Example

Let $X_1, \dots, X_n$ be independent **Bernoulli random variables** with parameter p .

Sums of random variables – moment generating function

The moment generating function can be used to compute moments of random variables.

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Then $M_{X_i}(s) = (1-p)e^{0s} + pe^{1s} = 1-p+pe^s$, $i = 1, \dots, n$.

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Z is again a Poisson random variable, this time with the parameter $\lambda_1 + \lambda_2$:

$$P(Z = k) = \frac{(\lambda_1 + \lambda_2)^k}{k!} e^{-(\lambda_1 + \lambda_2)}.$$

Compare with the difficulty of a direct computation of the convolution.

Summary

Joint distribution function of a random vector (X, Y) :

$$F_{X,Y}(x, y) = P(X \leq x \cap Y \leq y).$$

Discrete random variables X and Y

Continuous random variables X and Y

Joint probabilities of values / density:

$$P(X = x \cap Y = y)$$

$$f_{X,Y}(x, y)$$

Marginal probabilities / density of X :

$$P(X = x) = \sum_y P(X = x \cap Y = y)$$

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy$$

$$P(Y = y) = \sum_x P(X = x \cap Y = y)$$

$$f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dx$$

Independence of X and Y :

$$P(X = x \cap Y = y) = P(X = x) P(Y = y)$$

$$f_{X,Y}(x, y) = f_X(x) f_Y(y)$$

Covariance of X and Y :

$$\text{cov}(X, Y) = E[(X - E X)(Y - E Y)] = E[XY] - E X E Y$$

X and Y are called non-correlated whenever $\text{cov}(X, Y) = 0$.

If X and Y are independent, then they are also non-correlated.